

Statistical Reasoning

Week 12

Sciences Po - Louis de Charsonville

Spring 2018

Outline

Research Paper

Review of regression

Instrumental Variables

Review of Instructions

Research Paper

Timeline

Final draft 1st May

Review of regression

Regression models produce **fitted** (predicted) values and residuals that hold the unexplained variance for each data point. Issues that arise in that context are :

- ▶ unreliable coefficients due to **multicollinearity**, i.e. interactions between independent variables
- ▶ unreliable significance tests due to **heteroskedasticity**, i.e. heterogeneous variance in the residuals
- ▶ unreliable predictions due to **outliers and influential points** in the data that either do not fit or 'overfit' the model

Note : The model still assumes a linear, additive relationship between Y and X_1, X_2, \dots, X_k . That assumption can also be violated among other matters.

The model also **fits** a linear function to the data, of the form :

$$Y = \alpha + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k + \epsilon \quad (1)$$

where :

- ▶ Y is the **dependent variable** (response)
- ▶ X is a vector of **independent variable** (predictors)
- ▶ α is the constant
- ▶ $\beta_1 X_1 + \beta_2 X_2 + \dots + \beta_k X_k$ is a vector of regression coefficients
- ▶ ϵ is the **error term**

```
reg births schooling log_gdpc
```

The `reg` command can take any number of continuous variables as arguments, and shows **unstandardised** coefficients by default, using their original metric and possible transformation :

```
. reg births schooling log_gdpc
```

Source	SS	df	MS			
Model	150.301883	2	75.1509417	Number of obs =	86	
Residual	70.475313	83	.849100157	F(2, 83) =	88.51	
Total	220.777196	85	2.59737878	Prob > F =	0.0000	
				R-squared =	0.6808	
				Adj R-squared =	0.6731	
				Root MSE =	.92147	

births	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
schooling	-.1976117	.0724595	-2.73	0.008	-.3417306	-.0534927
log_gdpc	-.4703416	.1324501	-3.55	0.001	-.7337796	-.2069036
_cons	7.950304	.6861182	11.59	0.000	6.585642	9.314965


```
reg births schooling log_gdpc, beta
```

The beta option provides **standardised coefficients**, which use the standard deviation of regressors (or predictor, i.e. the independent variables) in order to provide coefficients with comparable units :

births	Coef.	Std. Err.	t	P> t	Beta
schooling	-.1976117	.0724595	-2.73	0.008	-.3686479
log_gdpc	-.4703416	.1324501	-3.55	0.001	-.4800156
_cons	7.950304	.6861182	11.59	0.000	.

```
reg births schooling i.region
```

Categorical variables can be used as **dummies**, i.e. binary recodes of each category that are tested against a reference category to provide regression coefficients for net effect of that category alone :

births	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
schooling	-.0415563	.0639718	-0.65	0.518	-.1688888	.0857763
log_gdpc	-.742187	.1380037	-5.38	0.000	-1.016876	-.4674975
region						
2	-.6523485	.5803126	-1.12	0.264	-1.807432	.5027349
3	.3682404	.254364	1.45	0.152	-.1380585	.8745393
4	1.411177	.2486027	5.68	0.000	.9163457	1.906008
5	1.167491	.337383	3.46	0.001	.4959471	1.839035
_cons	8.315004	.8006456	10.39	0.000	6.721359	9.908649

Instrumental Variables

- ▶ Some variables might be *unobserved*.
- ▶ OLS is inconsistent under omitted variables (Week 10).
- ▶ Omitted variables bias can be mitigated using **proxy variable** for the unobserved variable.
- ▶ Suitable proxy variable are not always available.
- ▶ When treatment is not randomly assigned, the causal effect of the treatment cannot be recovered from simple regression methods

Example

$$\log(wage) = \beta_0 + \beta_1 educ + \beta_2 ability + \epsilon \quad (2)$$

- ▶ *ability* is unobserved
- ▶ no proxy $\rightarrow \log(wage) = \beta_0 + \beta_1 educ + u$
- ▶ u contains *ability* and β_1 is biased if *educ* and *ability* are correlated.

Simple OLS model

$$\log(\text{wage}) = \beta_0 + \beta_1 \text{educ} + \epsilon \quad (3)$$

```
. reg lwage educ
```

Source	SS	df	MS	Number of obs	=	428
-----+-----				F(1, 426)	=	56.93
Model	26.3264237	1	26.3264237	Prob > F	=	0.0000
Residual	197.001028	426	.462443727	R-squared	=	0.1179
-----+-----				Adj R-squared	=	0.1158
Total	223.327451	427	.523015108	Root MSE	=	.68003

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
educ	.1086487	.0143998	7.55	0.000	.0803451	.1369523
_cons	-.1851969	.1852259	-1.00	0.318	-.5492674	.1788735
-----+-----						

- ▶ One additional year of education is associated with earnings 11% higher.
- ▶ **Bias** : Self-selection into education → individuals who have the most to gain from education are the most likely to stay.
- ▶ Ability is unobserved and is correlated with both education and wages.
- ▶ OLS estimates are not consistent.

Solutions

- ▶ One additional year of education is associated with earnings 11% higher.
- ▶ **Bias** : Self-selection into education → individuals who have the most to gain from education are the most likely to stay.
- ▶ Ability is unobserved and is correlated with both education and wages.
- ▶ OLS estimates are not consistent.

Solutions

- ▶ Randomized control trial (RCT) : allocate education randomly to individuals and observe the difference in their wages.
 - ▶ However : RCT is infeasible on ethical grounds.
- ▶ Quasi-natural experiments can alter individuals choices and can be used as instruments.

A **valid instrument** (or instrumental variable, IV) is :

1. Significantly correlated with the treatment of interest
(**instrument relevance**)
2. Only affect the outcome via its effect on the treatment
(exclusion restriction or **instrument exogeneity**)

Formally :

$$y = \alpha + \beta x + \epsilon \quad (4)$$

z is a valid instrument if :

1. Instrument relevance $\Leftrightarrow Cov(z, x) \neq 0$
2. Instrument exogeneity $\Leftrightarrow Cov(z, \epsilon) = 0$

While we can test whether the first condition is satisfied the second condition cannot be tested.

Examples of instruments ?

- ▶ IQ (Intelligence Quotient) ?
- ▶ Mother's education ?
- ▶ Number of siblings ?
- ▶ Legislative change increasing number of minimum schooling

Example 1 - Father's education

- ▶ Assume father's education is uncorrelated with ϵ
- ▶ We can check father's education is indeed correlated with education

```
reg educ fatheduc if !mi(lwage)
```

Source	SS	df	MS	Number of obs	=	428
-----+-----				F(1, 426)	=	88.84
Model	384.841983	1	384.841983	Prob > F	=	0.0000
Residual	1845.35428	426	4.33181756	R-squared	=	0.1726
-----+-----				Adj R-squared	=	0.1706
Total	2230.19626	427	5.22294206	Root MSE	=	2.0813

educ	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----						
fatheduc	.2694416	.0285863	9.43	0.000	.2132538	.3256295
_cons	10.23705	.2759363	37.10	0.000	9.694685	10.77942
-----+-----						

Example 1 - Father's education

- ▶ We use father's education as a IV for educ :

```
. ivreg lwage (educ = fatheduc)
```

```
Instrumental variables (2SLS) regression
```

Source	SS	df	MS	Number of obs	=	428
Model	20.8673618	1	20.8673618	F(1, 426)	=	2.84
Residual	202.460089	426	.475258426	Prob > F	=	0.0929
Total	223.327451	427	.523015108	R-squared	=	0.0934
				Adj R-squared	=	0.0913
				Root MSE	=	.68939

lwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
educ	.0591735	.0351418	1.68	0.093	-.0098994	.1282463
_cons	.4411035	.4461018	0.99	0.323	-.4357311	1.317938

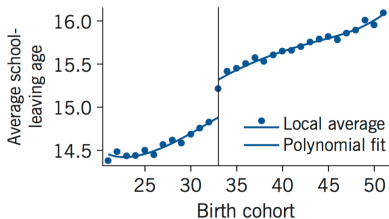
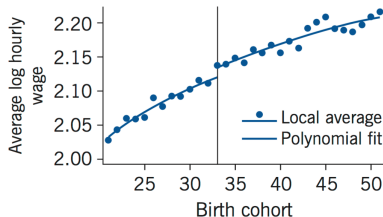
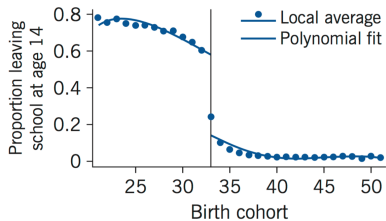
```
Instrumented: educ
```

```
Instruments: fatheduc
```

Example 2 - Legislative change in number of mandatory schooling

- ▶ In 1947, a legislative change in the UK increased the minimum school leaving age from 14 to 15
- ▶ Children who wanted to leave school at 14 are prevented from doing so and have to do one additional year of schooling.
- ▶ Let assume :
 - ▶ children under the two legislations are similar
 - ▶ Children face similar labor market conditions
- ▶ Quasi-natural experiment : independent of their ability, some individuals will need to stay one more year in schooling.
- ▶ Instrument variable : binary variable for being affected by the reform.

Example 2 - Legislative change in number of mandatory schooling



Example 2 - Legislative change in number of mandatory schooling

- ▶ Impact of the IV (the reform) on the treatment (education) (1st stage) :
 - ▶ Reform increased the average years of schooling for men by 0.397 years
- ▶ Impact of the IV on the dependent variable (wages) (Reduced-form estimate)
 - ▶ Reform increased wages by 1.2%
- ▶ IV estimates is $\frac{0.012}{0.937} = 0.03$ or 3% (Wald estimates).

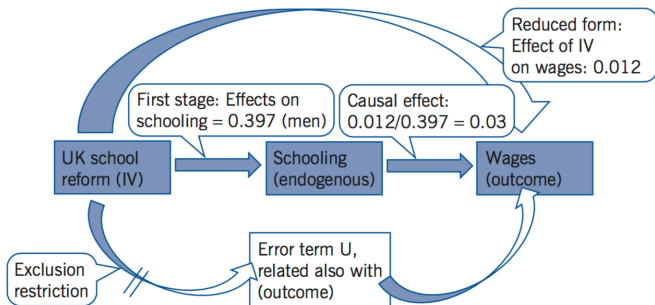
Example 2 - Legislative change in number of mandatory schooling

1. If the reform has an effect on education
2. If the reform affects wages *exclusively* through its effect on education

⇒ The IV estimates can be interpreted as the **causal effect of the treatment on the outcome.**

Schematic depiction of IV estimation

Example 2 - Legislative change in number of mandatory schooling



Note : A UK reform that increased minimum school leaving age is used as the Instrumental variable (IV) ; it should affect the outcome only via its effect on the endogenous variable but not in other ways

Credits : Sascha O. Becker, University of Warwick

- ▶ Causal relationship of interest :

$$Y = \alpha + \beta X + \epsilon$$

- ▶ First-Stage regression :

$$X = \eta + \gamma Z + u$$

- ▶ Second-Stage regression :

$$Y = \mu + \rho \hat{X} + v$$

- ▶ Reduced form :

$$Y = \delta + \phi Z + v$$

- ▶ **Wald estimate** is the ratio of the reduced form estimate and the first stage estimate
- ▶ Can be easily computed when the instrument takes only two values
- ▶ In general case, a "two stage least squares" (**2SLS**) estimate will be computed
- ▶ Only the variation in the treatment coming from the instrument is used to explain the variance in the outcome.

Difficulties

- ▶ Finding a valid instrument
- ▶ Interpreting the results

1. Relevance

- ▶ Correlation between the instrument and the change in treatment allocation is strong.
- ▶ **Weak instruments** = instruments that are only *weakly* correlated with the treatment.
- ▶ Weak instruments induce a bias that can be larger than the bias of the OLS estimates.

2. Exclusion restriction

- ▶ Cannot be statistically tested
- ▶ Need to be supported by a convincing narrative

Intepreting IV results can be difficult...

Why is the IV estimate much lower than OLS estimate?

Interpreting IV results can be difficult...

Why is the IV estimate much lower than OLS estimate ?

- ▶ OLS estimate describes the average difference in earnings for those whose education differs by one year
- ▶ IV estimate is the effect of increasing education *only* for the population whose choice of the treatment was *affected* by the instrument.
- ▶ Such effect is known as **Local Average Treatment Effect (LATE)**
- ▶ In this case, treatment effects are heterogeneous.
- ▶ For IV to estimate LATE, another assumption need to be satisfied :
 - ▶ While the instrument may have no effect on some people, all those who are affected are affected in the same way.

Monotonicity assumption

Some LATE's specific jargon :

- ▶ **Always-taker** : They always take the treatment independently of the IV.
- ▶ **Compliers** : Their treatment status is affected by the instrument in the right direction.
- ▶ **Never-takers** : They never take the treatment independently of IV.
- ▶ **Defiers** : Their treatment status is affected by the instrument in the "wrong" direction.

⇒ **Monotonicity ensures that there are no defiers.**

- ▶ With defiers, effects on compliers could be partly cancelled out by opposite effects on defiers
- ▶ Reduced form effect could be close to 0 although treatment effects are positive for everyone (but the compliers are pushed in one direction by the instrument and the defiers in the other direction)

Example

		Old regime	
		<i>Educ = 14</i>	<i>Educ ≥ 14</i>
New regime	<i>Educ = 14</i>	Never-taker	Defier
	<i>Educ ≥ 14</i>	Complier	Always-taker

IV - Wrapping up

Pros

- ▶ IV are useful to address :
 - ▶ Omitted variable bias
 - ▶ Measurement error
 - ▶ Simultaneity or reverse causality

Cons

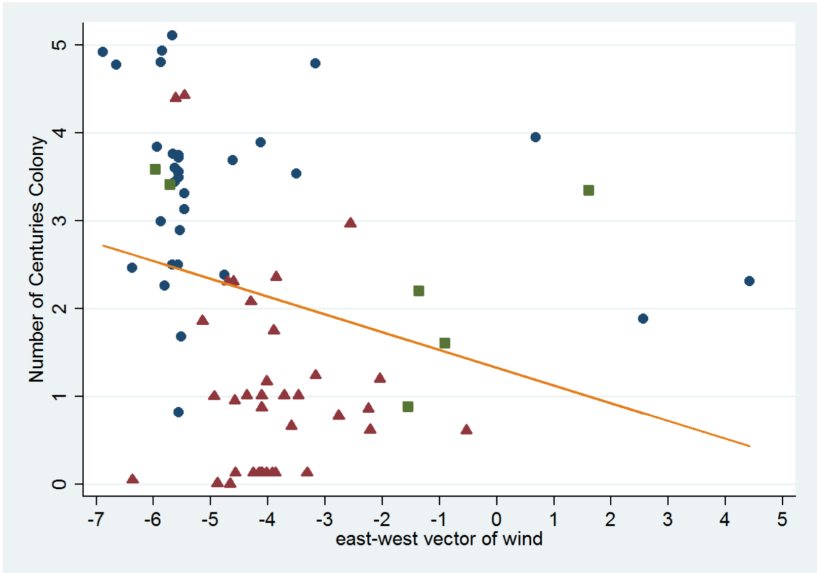
- ▶ Finding valid instrumental variables that affect treatment but do not have a direct effect on the outcome is difficult.
- ▶ Estimated treatment effects do not generally apply to the whole population
- ▶ Estimated treatment effects may vary across different instruments.
- ▶ In case of “weak” instruments, instrumental variable estimates are biased.

Institutions and prosperity

- ▶ In rich economies institutions (rules that govern society) function rather well on the whole while in poor ones they don't.
- ▶ Is good institutions a cause of economics progress or a consequence ?
- ▶ Find an IV which is link to institutions but not to economic success.

Institutions and prosperity

- ▶ In rich economies institutions (rules that govern society) function rather well on the whole while in poor ones they don't.
- ▶ Is good institutions a cause of economics progress or a consequence ?
- ▶ Find an IV which is link to institutions but not to economic success.
- ▶ Feyrer and Sacerdote (2006) uses **winds and currents** as an IV.
- ▶ Early colonists went where their sails took them. Some islands were colonized earier because there lay on natural sailing routes



Institutions and prosperity - Findings

- ▶ A robust positive relationship between the years of European colonialism and current levels of income : a century as a colony is worth a 40% increase in today's GDP.
- ▶ Years under US and Dutch colonial rule are significantly better than years under the Spanish and Portuguese.
- ▶ Later years of colonialism are associated with a much larger increase in modern GDP than years before 1700.

Review of Instructions

Univariate statistics

- ▶ Introduction
- ▶ Datasets
- ▶ Distribution
- ▶ Estimation

Bivariate statistics

- ▶ Significance
- ▶ Crosstabs
- ▶ Correlation
- ▶ Regression

Statistical modelling

- ▶ Basics
- ▶ Extensions
- ▶ Diagnostics
- ▶ Conclusion

Tablets of Stones

1. Interpret your results
2. Reference your sources
3. Proofread your work



Thank you

`exit, clear`

- ▶ Francois Briatte & Ivaylo Petev, Stata Guide
- ▶ Urdan, Statistics in Plain English
- ▶ Jeffrey M. Wooldridge, Introductory Econometrics : A Modern Approach, 5e Ed.
- ▶ Marcelo Coca Perrillon, Health Services Research Methods I, University of Colorado
- ▶ Michael Visser, Econometrics I, ENSAE ParisTech
- ▶ Sasha Becker, Using instrumental variables to establish causality
- ▶ Fabian Waldinger, Applied Econometrics, University of Warwick